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Research Article

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The Effect of Primary Surgical Technique for Treatment of Endometrial Cancer and Timing of Adjuvant Radiation Therapy

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Abstract

Objective: To determine if surgical approach (open, laparoscopic, robotic) influences time to initiation of adjuvant radiation therapy in women with endometrial cancer.

Methods: The National Cancer Database was used to search for patients with stage I to III endometrial cancer who received adjuvant radiation therapy from 2010-2012. Demographic, socioeconomic and clinical information was abstracted for each patient. Time to initiation of adjuvant radiation therapy was compared between groups using univariable (unadjusted) and multivariable (adjusted for all demographic and clinical variables) Cox regression models.

Results: A total of 15,480 patients were included in our study. 47.9% of patients underwent laparotomy or an unspecified surgical approach, 38.6% underwent robotic surgery, and 13.5% had a laparoscopic surgery. The distribution of time to radiation was significantly different among these groups with hazards ratios 1.0 (reference), 1.1 [95% CI 1.07, 1.15] and 1.15 [95% CI 1.02-1.06] for open, robotic and laparoscopic surgery (p<0.0001), respectively. After adjusting for covariates, younger age and worse disease status as measured by stage and grade are associated with longer wait times to radiation.

Conclusions: Women who underwent open staging surgery for endometrial cancer experienced delays from surgery to initiation of adjuvant radiation therapy as compared to women who had minimally invasive surgery.

Introduction

Endometrial cancer (EC) is the most common gynecologic malignancy in the United States, accounting for an estimated 63,230 new cases and 11,530 deaths in 2018 [1]. Women with endometrial cancer undergo surgical staging followed by adjuvant treatment as appropriate. Adjuvant radiotherapy (RT) is commonly used for patients with early-stage high-intermediate risk disease or those with advanced disease in order to prevent locoregional recurrence. The locoregional recurrence rate for early stage endometrial cancer with risk factors is up to 26% [2,3]. For patients with advanced stage disease, despite receiving systemic chemotherapy, up to 18% of patients will develop a locoregional recurrence [4].

When feasible, minimally invasive surgery is the preferred surgical modality for EC staging and treatment [5]. Laparoscopic surgical management for EC has been shown to be a safe alternative to laparotomy. Laparoscopy as compared to open techniques results in fewer short-term complications and decreased hospital length of stay [6] with similar oncological outcomes [5,7]. A meta-analysis of the literature comparing robotic-assisted laparoscopic surgery to laparotomy for EC also revealed decreased intraoperative and postoperative complication rates, decreased postoperative morbidity, improved patient-reported outcomes, including significantly shorter return to daily activities, and equivalent



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survival rates [8,9]. Robotic-assisted surgery for EC has also been compared with traditional laparoscopic approach and has shown equivalent short-term outcomes [10]. The use of laparoscopy and robotic-assisted surgery for EC is increasing.

Mode of surgery influences patient recovery times and therefore may influence time from surgery to adjuvant therapy. There is evidence that time intervals between diagnosis, surgery, and initiating adjuvant treatment for patients with endometrial cancer influences outcomes [11-13]. Studies have shown that a delay between hysterectomy and adjuvant RT might portend worse outcomes. Currently, literature exploring this relationship is limited. One retrospective study showed that initiating adjuvant RT more than six weeks after surgery decreased disease-specific survival for EC [11]. Others have found that local recurrence rate was associated with time interval from surgery to RT using a cutoff of nine weeks [12,13]. Given the advent of minimally invasive surgery, we sought to evaluate the impact of surgical technique on the time interval for initiating adjuvant radiation therapy for patients with endometrial cancer.

Methods

The National Cancer Database (NCDB) is a national hospital-based cancer registry that is a joint endeavor of the American College of Surgeons and American Cancer Society. Annually, data from over 1 million patients representing 70% of all new cancer diagnoses in the United States are reported to the NCDB. Approximately 1,500 Commission on Cancer (CoC)-accredited hospitals contribute deidentified data. Data reporting is highly standardized to the CoC Registry Manuals, the American Joint Committee on Cancer and Collaborative Stage manuals, and the International Classification of Diseases for Oncology, Third Edition (ICD-0-1) [14,15]. The Institutional Review Board at the Icahn School of Mount Sinai granted exemption status for this study.

The NCDB Participant User File was used to search for patients with stage I to III EC who underwent primary surgical treatment and for whom the surgical technique was known from 2010 to 2012. Only those patients with a known cancer diagnosis prior to surgery who also received adjuvant radiotherapy were included in the study. The types of adjuvant RT delivered were external beam radiation therapy (EBRT), brachytherapy, radioisotopes, or combined EBRT with brachytherapy boost or radioisotopes. Patients who received radiation outside of the uterus, cervix, pelvis or abdomen were

excluded. Demographic and socioeconomic information was abstracted for each patient, including age, year of diagnosis, race/ethnicity, income, insurance status, and facility location. Clinical data collected included Charlson-Deyo comorbidity score, surgical approach, time to radiation, and time hospitalized. Pathologic details recorded included International Federation of Gynecology and Obstetrics (FIGO) staging and grade of disease.

Patients were categorized into three groups according to surgical approach: robotic, laparoscopic, and open or unspecified. Patients whose surgical approach was not reported were categorized as unspecified and grouped with patients who underwent open surgery by the NCDB. Demographic and clinical characteristics were compared between the groups using one-way ANOVA or Kruskal-Wallis tests for continuous measures and Chi-square tests for categorical measures. The distribution of time to adjuvant radiation by surgical approach was estimated using the Kaplan Meier method and differences between the groups were assessed using a log-rank test. Univariable (unadjusted) and multivariable (adjusted for all demographic and clinical variables) Cox regression models were estimated to assess each factor's association with time to adjuvant radiation therapy. All analyses were conducted using SAS version 9.4 (SAS, Cary, NC).

Results

A total of 15,480 patients diagnosed with endometrial cancer from 2010 to 2012 met inclusion criteria. Of these, 47.9% had a laparotomy or unspecified surgical approach, 38.6% underwent robotic surgery, and 13.5% had a laparoscopic surgery. Although statistically significant due to the large sample size, there were no clinically significant differences between the groups based on age, income, insurance status, facility location, and Charlson-Deyo score (Table 1). The robotic group had a higher proportion of white patients compared to both the laparoscopic and open/unspecified groups. The open/unspecified group had proportionately more stage 3 and grade 3 and 4 (undifferentiated or anaplastic) patients compared to the robotic and laparoscopic groups. As expected, the length of stay in hospital was also significantly longer for open/ unspecified patients compared to patients who underwent less invasive procedures. Notably, the distribution of year of diagnosis varied by surgical group with the percentage of patients in both the robotic and laparoscopic increasing by year versus decreasing in the open/unspecified group.

Table 1: Characteristics of endometrial cancer patients treated with radiation therapy by surgical approach.

Variable	Robotic		Laparoscoj	oic	Open or Unspecified		P-Value*	
	N	%	N	%	N	%		
Total	5975	38.6	2096	13.5	7409	47.9		
	Age (Years)**							
Mean (standard deviation)	63.8 (10.1)		64.3 (10.1)		62.7 (10.5)			
Median (interquartile range)	64 (57, 71)		64 (58, 71)		63 (56, 70)			
	Year of Diagnosis							
2010	1403	23.5	569	27.2	2824	38.1		
2011	2040	34.1	714	34.1	2460	33.2		
2012	2532	42.4	813	38.8	2125	28.9		

		Ethnicity					<0.00	
Non-Hispanic White	5044	85	1646	79.2	5568	75.8		
Non-Hispanic Black	474	8	233	11.2	1082	14.7		
Non-Hispanic Asian	145	2.4	53	2.6	202	2.8		
Hispanic	226	3.8	122	5.9	436	5.9		
Other	43	0.7	25	1.2	60	0.8		
Median Income of Area of Residence								
< \$38,000	800	13.5	315	15.1	1331	18.1		
\$38,000-\$47,999	1339	22.5	410	19.6	1721	23.4		
\$48,000-\$62,999	1718	28.9	533	25.6	1894	25.7		
\$63,000 +	2090	35.1	827	39.7	2416	32.8		
	Insu	rance				'	<0.00	
Not Insured	178	3	63	3	415	5.8		
Private/Managed Care	2942	49.6	1010	48.6	3287	45.9		
Medicaid/Medicare/Government	2809	47.5	1004	48.3	3458	48.3		
	Facility	Location	ı				<0.00	
New England	440	7.4	183	8.7	491	6.6		
Middle Atlantic	1158	19.4	460	21.9	1489	20.1		
South Atlantic	1157	19.4	557	26.6	1273	17.2		
East North Central	1211	20.3	252	12	1701	23		
East South Central	352	5.9	58	2.8	330	4.5		
West North Central	561	9.4	191	9.1	627	8.5		
West South Central	241	4	95	4.5	586	7.9		
Mountain	283	4.7	109	5.2	251	3.4		
Pacific	572	9.6	191	9.1	661	8.9		
Charlson-Deyo Score								
0	4373	73.2	1540	73.5	5341	72.1	<0.00	
1	1364	22.8	454	21.7	1639	22.1		
2	238	4	102	4.9	429	5.8		
	Time Hospita			1.7	127	5.0	<0.00	
Mean (standard deviation)	1.55 (3.26)	lizeu (zuyo	1.98 (2.86)		3.84 (4.57)			
Median (interquartile range)	1 (1, 2)		1 (1, 2)		3 (2, 4)			
(4a-a-a-a-a-a-a-a-a-a-a-a-a-a-a-		age	- (*) =)		- (- , -)		<0.00	
1	4046	67.7	1448	69.1	4307	58.1	-0.00	
2	693	11.6	227	10.8	1118	15.1		
3	1236	20.7	421	20.1	1984	26.8		
3		ade	741	20.1	1701	20.0	<0.00	
1 1183 23.7 467 25.8 1188 19							\0.0t	
2	1870	37.4	649	35.9	2120	33.9		
3	1662	33.3		32.4		40.6		
ა	1002	33.3	586	32.4	2535	40.6		

^{*}Chi-square test unless otherwise noted

Across the three surgical groups, the distribution of time to radiation was significantly different (p<0.0001, Figure 1). Although median times to radiation are similar across the groups, the 75th percentile of the open or unspecified group is 108 days compared to 98 and 92 days for the robotic and laparoscopic groups respectively, meaning that more open patients had a longer wait time to radiation compared to robotic and laparoscopic patients (Table 2).

In unadjusted analyses, compared to open/unspecified patients, at any given time laparoscopic patients were 1.15 times more likely to initiate radiation therapy (95% CI 1.10,1.21) and robotic patients were 1.11 times more likely to initiate radiation therapy (95%CI 1.07, 1.15; Table 3). Age, race/ethnicity, facility location, time hospitalized, stage and grade were also significantly associated with time to radiation therapy in univariable analyses.

^{**}P-value is from a one-way ANOVA

^{***}P-value is from a Kruskal-Wallis test

In a multivariable model, adjusting for these factors as well as year of diagnosis, income, insurance status, and Charlson-Deyo score, surgical approach overall did not retain statistical significance (p=0.07); however, laparoscopic approach remained associated with quicker times to radiation (robotic HR 1.00 [95% CI 0.96, 1.04]; laparoscopic HR 1.06 [95% CI 1.01, 1.13]). Age, facility

location, time hospitalized, stage, and grade remained statistically significant after adjustment (Table 3). After adjusting for all other covariates, being younger and having worse disease status as measured by stage and grade are associated with longer wait times to radiation.

Table 2: Time to radiation therapy percentiles for endometrial cancer patients by surgical approach.

Time to Radiation (Days)	Surgical Approach							
	Robotic	Laparoscopic	Open or Unspecified					
25 th Percentile (95% CI)	47 (46,48)	46 (45,47)	46 (45,47)					
50 th Percentile (95% CI)	63 (62,64)	60 (58,62)	63 (62,64)					
75 th Percentile (95% CI)	98 (96,100)	92 (89,97)	108 (106,111)					

Table 3: Unadjusted and adjusted hazard ratio estimates.

Variable	Un	Univariable P-Value		P-Value*	Multivariable			P-Value*
	HR	95%	CI		HR	95%	CI	
Surgical Appro	oach			<0.001				0.07
Open or Unspecified		1.0 (ref)			1.0 (ref)			
Robotic	1.11	1.07	1.15		1	0.96	1.04	
Laparoscopic	1.15	1.1	1.21		1.06	1.01	1.13	
Age (10 Years)	1.04	1.02	1.06	<0.001	1.04	1.02	1.07	< 0.001
Year of Diagno	osis			0.07				0.15
2010		1.0	0 (ref)		1.0 (ref)			
2011	1	0.96	1.04		0.98	0.93	1.02	
2012	1.04	1	1.08		1.02	0.97	1.07	
Race/Ethnic	ity			<0.001				0.61
Non-Hispanic White		1.0	0 (ref)			1.0	(ref)	
Non-Hispanic Black	0.86	0.82	0.9		0.96	0.9	1.02	
Non-Hispanic Asian	0.8	0.73	0.89		0.98	0.87	1.1	
Hispanic	0.93	0.87	1		1.02	0.94	1.11	
Other	1.09	0.92	1.3		1.05	0.85	1.29	
Median Income of Area of Residence			0.95	0.95				
<\$38,000	1.0 (ref)				1.0 (ref)			
\$38,000-\$47,999	1	0.95	1.05		1.01	0.95	1.07	
\$48,000-\$62,999	0.99	0.94	1.04		1	0.94	1.06	
\$63,000 +	0.99	0.95	1.04		1.01	0.95	1.07	
Insurance				0.50	0.50			
Not Insured		1.0	0 (ref)		1.0 (ref)			
Private/Managed Care	1.01	0.93	1.09		1.04	0.94	1.14	
Medicaid/Medicare/Government	1.03	0.95	1.11		1.01	0.92	1.12	
Facility Locat	ion			<0.001	<0.001			
New England		1.0	0 (ref)	1.0 (ref)				
Middle Atlantic	0.87	0.82	0.93		0.91	0.84	0.98	
South Atlantic	0.89	0.83	0.96		0.98	0.91	1.06	
East North Central	1.01	0.94	1.08		1.04	0.96	1.12	
East South Central	0.96	0.88	1.06		1.11	0.99	1.23	
West North Central	1	0.93	1.08		1.07	0.98	1.17	
West South Central	1.15	1.05	1.25		0.93	0.84	1.04	
Mountain	0.92	0.84	1.01		1.02	0.91	1.14	

Pacific	0.67	0.62	0.72		0.74	0.68	0.81	
Charlson-Deyo Score			0.84				0.54	
0	1.0 (ref)				1.0 (ref)			
1	1	0.96	1.04		0.98	0.94	1.03	
2	1.02	0.95	1.1		1.03	0.95	1.12	
Time Hospitalized (Days)	0.98	0.98	0.99	<0.001	0.99	0.98	0.99	<0.001
Stage			<0.001	<0.001				
1	1.0 (ref)					1.0 (ref)		
2	0.94	0.9	0.99		0.95	0.89	1	
3	0.52	0.5	0.55		0.51	0.49	0.53	
Grade			<0.001	<0.001				
1	1.0 (ref)				1.0 (ref)			
2	0.94	0.89	0.98		0.95	0.9	0.99	
3	0.69	0.65	0.72		0.68	0.64	0.71	
4	0.61	0.56	0.66		0.61	0.56	0.66	

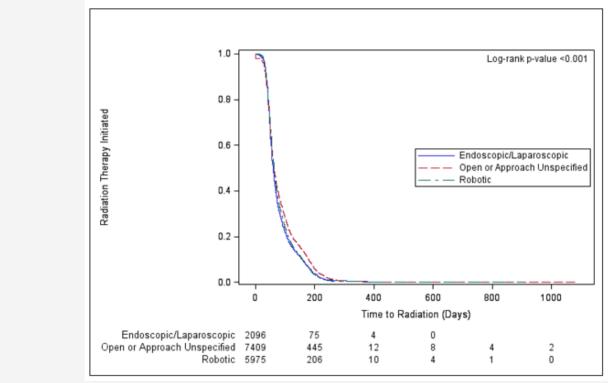


Figure 1: Kaplan Meier estimates of time to radiation therapy for endometrial cancer patients by surgical approach.

Discussion

To the best of our knowledge, this is the first report in a large database to have studied the impact of surgical technique on time to initiation of adjuvant RT in patients with endometrial cancer. The study's main findings are that younger age, later stage at diagnosis, higher grade of tumor at diagnosis and laparotomy were all associated with longer surgery-to-radiation intervals. Laparotomy resulted in delays to initiation of RT when compared with robotic-assisted laparoscopy and conventional laparoscopy.

The majority of the existing literature affirming the clinical significance of delays in adjuvant RT after surgery is in patients with locally advanced breast cancer. The largest and most recent meta-

analysis evaluating the effect of waiting times for postoperative RT on breast cancer patient outcomes included a review of 34 publications encompassing a total of 79,616 patients [16]. This study reported that delays in post-lumpectomy adjuvant RT are significantly associated with risk of local recurrence in patients with locally advanced disease [16].

Within the gynecologic oncology literature, most of the studies evaluating the clinical impact of time to adjuvant therapy after surgery are in patients with cervical cancer. A retrospective study performed in Thailand identified 125 patients over 29 years who underwent radical hysterectomy for stage IA2 or IB1 squamous cervical cancer who received either adjuvant radiation therapy

or adjuvant concurrent chemoradiation [17]. The investigators concluded that a time interval greater than 4 weeks between surgery and adjuvant therapy was associated with a worse recurrence-free survival [17]. Jhawar S, et al [18] performed a large-scale multi-institutional analysis of 3051 patients with early cervical cancer undergoing radical hysterectomy followed by adjuvant chemoradiation using the National Cancer Database [18]. The results of this study indicate that a delay of greater than 7 weeks from surgery to initiation of chemoradiation is associated with worse survival outcomes [18]. In a retrospective study of 226 patients with early stage cervical cancer, You KY, et al [19] similarly investigated the clinical significance of interval from surgery to adjuvant therapy by first dividing patients into two cohorts by FIGO staging - IA2-IB1 and IB2-IIA [10]. They concluded that for patients with stage IB2-IIA disease, overall survival and disease-free survival were higher for patients who had an interval to adjuvant treatment less than 5 weeks, but no similar effect was observed for patients with earlier stage disease [19].

Very few studies specifically evaluate differences in clinical outcome due to delay to adjuvant radiation therapy in endometrial cancer patients. In a retrospective single-institution study by Ahmad et al in 1995, investigators reviewed charts for 195 endometrial cancer patients collected over 22 years who all received adjuvant whole pelvic external beam radiation therapy following surgical staging [11]. Based on their findings, they concluded that a surgeryto-radiation delay of 6 weeks or greater was significantly associated with decreased disease specific survival [11]. This study, however, did not specifically evaluate surgical approach as a potential factor accounting for delays in treatment. A similar retrospective singleinstitution study by Cattaneo, et al in 2016 reviewed 308 patient charts over 22 year period which showed a surgery-to-radiation delay of 9 weeks or greater to be an independent predictor of recurrence as well as worse disease specific survival and overall survival in patients with endometrial cancer [13].

To our knowledge only one study specifically explores surgical approach as a possible factor responsible for delays to initiation of adjuvant RT in endometrial cancer patients. In a retrospective single-institution study, Stahl et al compared 74 patients who underwent robotic-assisted laparoscopic hysterectomy for staging for stage I-II endometrial carcinoma with 63 patients who underwent total abdominal hysterectomy for the same indication [20]. They concluded that while the average time interval from surgery to radiation oncology consultation was similar between the two groups (30 days), that patients who underwent robotic-assisted laparoscopic hysterectomy had on average a one week delay in actual initiation of treatment due to longer vaginal cuff healing time [20].

In direct comparison, our study found that laparotomy leads to longer interval time to initiation of adjuvant therapy as compared to a minimally invasive approach. Possible reasons for delays to initiation of adjuvant RT in this cohort could include longer hospital stays, longer recovery time and increased risk of post-

operative complications in patients undergoing laparotomy versus minimally invasive procedures. It is also possible that patients with worse pre-operative disease burden (higher grade and later stage disease) were more likely to be selected for laparotomy than minimally invasive approaches and therefore possibly also had worse pre-operative functional status which could contribute to longer recovery time and delay to initiation of adjuvant therapy.

The strengths of our study include utilization of a large-scale national database, which allowed for a much larger sample size. Additionally, because our cohort is multi-institutional our data is more generalizable when compared to single institution studies where there may be site-specific confounding variables with regard to surgical technique, clinical care, or follow-up that could affect results.

There are several limitations to our study that are inherent to the use of a national, hospital-based registry. Prior to 2010, information regarding mode of surgery was not recorded, so data for this study was limited to after 2010. Furthermore, the NCDB does not describe reasons why adjuvant RT may have been delayed nor does it provide specific information regarding adjuvant treatment regimens (i.e. whether adjuvant chemotherapy was needed, and adjuvant radiation was intentionally planned for after completion of chemotherapy). Those patients with advanced disease or earlystage high-risk histology that received adjuvant chemotherapy in addition to adjuvant RT might have in fact started their adjuvant treatment earlier than the results seen. Lastly, because we did not have information regarding initial treatment plans for these patients, we could not assess if there were patients in our cohort who were supposed to receive RT but were lost to follow-up, changed providers, or passed away before treatment had begun. As is the case with other published studies utilizing NCDB, we recognize that there are limitations with using this data and hence the conclusions drawn from this study warrant additional research and investigation.

Thus, younger age, higher stage or grade at diagnosis, and laparotomy all contribute to longer delays to initiation of adjuvant radiation therapy after surgery. While prospective studies are needed to better assess how delays in initiation of adjuvant RT affect clinical outcomes, these studies would be difficult to perform given our desire to have every patient treated as expeditiously as possible. As such, efforts should be made by individual institutions by means of quality improvement studies to identify potential site-specific causes accounting for delays to initiation of adjuvant RT. Subsequently, attention should be given to addressing identified barriers to receiving care and ensuring streamlined and timely follow-up for all patients. Gynecologic oncologists and radiation oncologists are encouraged to remain cognizant of the timing between surgery and adjuvant RT for local control and initiate therapy at the soonest, safest time point.

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None.

Conflict of Interest

Authors declare no conflict of interest.

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