



# Appendix 1

## The Gauss-Markov Theorem

Suppose that the assumptions made in hold and that the errors are homoscedastic. The OLS estimator is the best (in the sense of smallest variance) linear conditionally unbiased estimator (BLUE) in this setting.

Let us have a closer look at what this means:

- Estimators of  $\beta_1$  that are linear functions of the  $Y_1, \dots, Y_n$  and that are unbiased conditionally on the regressor  $X_1, \dots, X_n$  can be written as  $\tilde{\beta}_1 = \sum_{i=1}^n a_i Y_i$  where the  $a_i$  are weights that are allowed to depend on the  $X_i$  but not on the  $Y_i$ .
- We already know that  $\tilde{\beta}_1$  has a sampling distribution:  $\tilde{\beta}_1$  is a linear function of the  $Y_i$  which are random variables. If  $E(\tilde{\beta}_1 | X_1, \dots, X_n) = \beta_1$ ,  $\tilde{\beta}_1$  is a linear unbiased estimator of  $\beta_1$ , conditionally on the  $X_1, \dots, X_n$ .
- We may ask if  $\tilde{\beta}_1$  is also the best estimator in this class, i.e., the most efficient one of all linear conditionally unbiased estimators where "most efficient" means smallest variance. The weights  $a_i$  play an important role here and it turns out that OLS uses just the right weights to have the BLUE property.

Consider the case of a regression of  $Y_1, \dots, Y_n$  only on a constant. Here, the  $Y_i$  are assumed to be a random sample from a population with mean  $\mu$  and variance  $\sigma^2$ . The OLS estimator in this model is simply the sample mean  $\hat{\beta}_1 = \frac{1}{n} \sum_{i=1}^n Y_i$ . Clearly, each observation is weighted by

$a_i = \frac{1}{n}$ . We also know that  $Var(\hat{\beta}_1) = \frac{\sigma^2}{n}$ . We now can use R to conduct a simulation study that demonstrates what happens to the variance of (5.3) if different weights  $w_i = \frac{1}{n} \pm \epsilon$  are assigned to either half of the sample  $Y_1, \dots, Y_n$  instead of using  $\frac{1}{n}$ , the OLS weights.

```
# set sample size and number of repetitions
n <- 100
reps <- 1e5

# choose epsilon and create a vector of weights as defined above
epsilon <- 0.8
w <- c(rep((1 + epsilon) / n, n / 2),
      rep((1 - epsilon) / n, n / 2))

# draw a random sample y_1,...,y_n from the standard normal distribution,
# use both estimators 1e5 times and store the result in the vectors 'ols' and
# 'weightedestimator'
ols <- rep(NA, reps)
weighted estimator <- rep(NA, reps)

for (i in 1: reps) { y <- rnorm(n)
  ols[i] <- mean(y)
  weighted estimator[i] <- cross prod (w, y) }
```



```
# plot kernel density estimates of the estimators' distributions:
```

```
# OLS
```

```
plot(density(ols),
```

```
  col = "purple",
```

```
  lwd = 3,
```

```
  main = "Density of OLS and Weighted Estimator",
```

```
  xlab = "Estimates")
```

```
# weighted
```

```
lines(density(weightedestimator),
```

```
  col = "steelblue",
```

```
  lwd = 3)
```

```
# add a dashed line at 0 and add a legend to the plot
```

```
Alline (v = 0, lty = 2)
```

```
legend ('top right',
```

```
  c ("OLS", "Weighted"),
```

```
  col = c ("purple", "steel blue"),
```

```
  lwd = 3)
```

---